

Course: CS5961/6951
Instructor: R. F. Riesenfeld
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Computational Statistics

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Notes: *Poisson Distribution as the Limit of Binomial Distribution*

Consider a small (Δt) interval in a larger interval $I = [I_a, I_b]$ of the real line, so that $I = N(\Delta t)$, for some large $N > 0$. In other words, $(\Delta t) = \frac{I}{N}$, so there are N intervals in I .

Now, suppose that $P\{x \in (\Delta t)_i\} = \lambda(\Delta t)_i$, where λ is a normalizing constant $(= \frac{1}{I})$. We can think of λ as linear probability density function, analogous to computing the total resistance of a wire as ρL , where ρ is a constant giving unit length resistance and L is the overall length of the wire under consideration.

From the Binomial Probability Distribution, we know how to compute the likelihood of k successes in N Bernoulli trials. Specifically,

$P\{k \text{ successes in } N \text{ trials}\} = \binom{N}{k} q^{N-k} p^k$, where p is the probability of success in a Bernoulli trial and q is the probability of failure. However, here we know that $p = \frac{\lambda I}{N}$ and $q = 1 - p$. Hence,

$$\begin{aligned}
 P\{k \text{ successes in } N \text{ trials}\} &= \frac{N!}{(N-k)!k!} \left(1 - \frac{\lambda I}{N}\right)^{N-k} \frac{\lambda I^k}{N^k} \\
 &= \frac{(\lambda I)^k}{k!} N(N-1) \cdot \dots \cdot (N-k+1) \left(1 - \frac{\lambda I}{N}\right)^{N-k} \left(\frac{1}{N}\right)^k
 \end{aligned}$$

At this point, noting that λ, I and k are fixed, we take the limit of the above equation as $N \rightarrow \infty$. Thus,

$$\begin{aligned} & \lim_{N \rightarrow \infty} P\{k \text{ successes in } N \text{ trials}\} \\ &= \lim_{N \rightarrow \infty} \left(\frac{(\lambda I)^k}{k!} N(N-1) \cdot \dots \cdot (N-k-1) \left(1 - \frac{\lambda I}{N}\right)^{N-k} \left(\frac{1}{N}\right)^k \right) \\ &= \frac{(\lambda I)^k}{k!} \lim_{N \rightarrow \infty} \left(N(N-1) \cdot \dots \cdot (N-k-1) \left(1 - \frac{\lambda I}{N}\right)^{N-k} \left(\frac{1}{N}\right)^k \right). \end{aligned}$$

And, with some slight rearranging, the following is arrived at,

$$\begin{aligned} &= \frac{(\lambda I)^k}{k!} \lim_{N \rightarrow \infty} \left(\frac{N(N-1)}{N} \cdot \dots \cdot \frac{(N-k-1)}{N} \left(1 - \frac{\lambda I}{N}\right)^{N-k} \right) \\ &= \frac{(\lambda I)^k}{k!} \lim_{N \rightarrow \infty} \left(\frac{N(N-1)}{N} \cdot \dots \cdot \frac{(N-k-1)}{N} \left(1 - \frac{\lambda I}{N}\right)^N \left(1 - \frac{\lambda I}{N}\right)^{-k} \right). \end{aligned}$$

First, we make the following trivial observations about the two limits,

$$\lim_{N \rightarrow \infty} \left(\frac{N(N-1)}{N} \cdot \dots \cdot \frac{(N-k-1)}{N} \right) = 1$$

and,

$$\lim_{N \rightarrow \infty} \left(1 - \frac{\lambda I}{N}\right)^N = 1.$$

Therefore, the subject limit reduces to the simpler expression,

$$= \frac{(\lambda I)^k}{k!} \lim_{N \rightarrow \infty} \left(1 - \frac{\lambda I}{N}\right)^N .$$

But, by the definition e given in Calculus, namely, $e^x = \lim_{n \rightarrow \infty} \left(1 + \frac{x}{n}\right)^n$, we have only to observe that the desired limit immediately becomes,

$$= \frac{(\lambda I)^k}{k!} e^{-\lambda I} .$$

Substituting the variable t for the arbitrary interval t leads us to the more common equivalent expression,

$$= \frac{(\lambda t)^k}{k!} e^{-\lambda t} .$$

The derived expression is known as the *Poisson Probability Distribution Function* for k events, given λ as the mean number successes in unit measure (i.e., time, area, or the like).

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